Martha J. Banks, The University of Chicago

In the last decade or so, telephone survey techniques have come to be considered seriously when designing a high quality survey. Reasons for this include the rising cost of field work (especially travel costs), a trend toward lower response rates in personal interview surveys (often due to respondents' fears of allowing strangers into their homes), and the recognition that most population groups have a fairly high phone coverage rate.

However, when examining data from a telephone survey, it would be extremely useful to have some idea of how estimates might differ from those which would have been obtained in a personal interview survey.

This study presents the results of one approach to measuring such differences. It uses data from a personal interview survey and estimates differences between the phone and the nonphone populations. Therefore it does not indicate all differences between results from a phone survey and a personal interview survey because it does not measure such things as differences in the answers given by the phone population when it is interviewed by phone rather than in person. It would require a special methodological study to obtain estimates of these differences. My analysis of existing data is a cost-effective way to approximate the differences between phone survey estimates and estimates from interviewing in person.

The data used in this analysis are from a national U.S. personal interview survey conducted for the Center for Health Administration Studies in 1976. This study focused on access to medical care in the United States. As part of this effort, black southerners living outsie of SMSAs and those of hispanic heritage living in the Southwest were oversampled at about 3.4 to 1 . Altogether, 7787 persons in 5432 families were interviewed. The overall response rate of this area probability survey was 85 percent.

Because the focus of the study was access to medical care, I will be presenting differences in medical care estimates between the phone and nonphone populations. However, I urge those who have access to data from other subject areas to conduct the same sort of examination on those data. I believe the results in the medical care area to be quite interesting, and it would be extremely useful to see how the results compare with those obtained in other subject areas.

During the interviewing, respondents were asked for their telephone numbers so that some of the interviews could be verified. The respondent also indicated where the phone was located. Therefore, we were able to construct a variable which identifies whether on not the interviewed persons had home phones.

Table 1 shows that about $10.1 \%$ of all families and $9.3 \%$ of all persons in the U.S. had no home phone in 1976. The best predictor of phone coverage is financial status, as measured by the last two variables shown in the table: family income and poverty status. Besides the low-income population, groups with low phone coverage are
southerners, especially rural southern blacks; hispanics; persons whose family head was under 25 and or divorced, separated, or never married; and those living alone or in large families of 7 or more.

Many people assume that coverage rates can be used as proxies for noncoverage bias measures. That is, they assume that the larger the percent without phones, the larger the differences between the total population and the phone population. However, data in this paper suggest that phone coverage is not as good a predictor of nonphone bias as commonly supposed.

Table 2 shows the effect on the estimated percent who have contacted a doctor during the preceeding year. Based on all persons, $76.7 \%$ of the population contacted a doctor. The figure for those with home telephones is $77.6 \%$. The ratio of the two, .988 (given in the last column), is significantly different from 1.000 at the five standard error level. Examining this last column of Table 2 shows that there are no population groups given for which the ratio of the total estimate to the phone estimate is significantly greater than 1.000 . All the ratios either are about 1.000 or are significantly below it. Therefore, using data from only the phone population would tend to overstate the percent seeing or speaking with a physician during the year.

On the other hand, comparing population subgroups using data for only those with phones would result in conclusions nearly identical to those based on comparing population groups using data for all persons. Both show that those in the Northeast are most apt to contact a doctor and those in the South (especially rural southern blacks) are the least likely to. Both data sets indicate that those in SMSAs are more likely to see or talk to a doctor than are those living outside SMSAS, as are preschool children and the divorced. Contacting a doctor is positively correlated with the finanical status of the family, as both the phone data and the total data show. Therefore, while a data set based on only the phone population may overstate the percent contacting a doctor within the year, estimates of differences between population subgroups may contain little bias.

Notice also that there is not a consistant relationship between phone coverage rates and the ratios between estimates. For example, Table 1 showed that persons in families whose head is under 25 have quite low phone coverage, only about $74 \%$. However in Table 2, the ratio between the total and phone population estimates, .995, certainly does not suggest a larger noncoverage bias for this group than for persons in families whose heads are 25 or older.

Other tables, not shown here, present the same conclusions for several other health care variables. Each indicates that the phone population is consistantly somewhat more health-care advantaged than is the total population.

There are several ways in which these results might be used. When analyzing telephone data, a researcher might merely keep in mind the fact that


[^0]TABLE 2 - PERCENT CONTACTING A DOCTOR DURING THE YEAR, BY PHONE COVERAGE; CHAS 1976*

|  | PERCENT C |  | CONTACTING A | DOCTOR | DURING THE YEAR |  | RATIUS OF THE FERCENTS |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| CHARACTERISTIC | $\begin{gathered} \text { PHONE } \\ \text { POPULATION } \end{gathered}$ |  | NONPHONE POPULATION |  | TOTAL POPULATION |  | NONPHONE TO PHONE | POPULATION POPULATICN | TOTAL POPULATION |  |
|  |  |  | TO PHONE | POPULATION |  |  |  |  |
| REGION |  |  |  |  |  |  |  |  |  |  |
| NORTHEAST | 81. 4\% | ( 1.4 ) |  |  | 74.8\% | ( 6.4) | $81.1 \%$ | (1.4) | 918 | (.081) | . 996 | (.004) |
| NORTH CENTRAL | 76.6 | (1.3) | 83.5 | ( 5.3) | 76.9 | (1.3) | 1.089 | (.071) | 1.004 | (.003) |
| SOUTH | 74.8 | (1.2) | 66.4 | ( 2.5) | 73.4 | (1.1) | . 888 | (.036) | . 981 | (.006) |
| WEST | 79.2 | (1.5) | 52.0 | ( 4.7 ) | 76.5 | (1.5) | . 656 | (.060) | 966 | (.006) |
| RESIDENCE |  |  |  |  |  |  |  |  |  |  |
| SMSA CENTRAL CITY | 78. 1 | (1.3) | 69.1 | ( 3.8 ) | 77.2 | (1.3) | . 885 | (.051) | 989 | (.005) |
| SMSA OTHER | 79.9 | (1.1) | 67.5 | (4.2) | 79:1 | (1.1) | . 845 | (.054) | . 990 | (.004) |
| NONSMSA URBAN | 75.1 | (1.9) | 68.7 | (4.2) | 74.4 | (1.7) | . 915 | (.061) | . 990 | (.007) |
| RURAL NONFARM | 76.0 | (1.6) | 66.5 | ( 3.7) | 74.7 | (1.5) | . 875 | (.052) | . 983 | (.007) |
| RURAL FARM | 70.6 | (3.2) | 63.3 | (10.6) | 70.2 | (3.1) | . 896 | (. 155 ) | . 994 | (.009) |
| RACE |  |  |  |  |  |  |  |  |  |  |
| SPANISH HERITAGE, SOUTHWEST | 72.0 | (5.7) | 44.0 | (11.0) | 64.7 | (5.3) | . 612 | (.160) | . 898 | (.042) |
| OTHEF WIIITE | 78.0 | (0.8) | 72.0 | ( 2.9 ) | 77.5 | (0.8) | . 924 | (.039) | . 995 | (.003) |
| NONSMSA SOUTHERN BLACK | 69.6 | (4.7) | 58.2 | (6.5) | 65.1 | (3.8) | . 827 | (. 110 ) | . 936 | (.043) |
| OTHER NONWHITE | 77.5 | (3.8) | 74.4 | ( 8.4) | 77.0 | (3.4) | . 960 | (.118) | . 994 | (.016) |
| AGE |  |  |  |  |  |  |  |  |  |  |
| O-5 | 89.3 | (1.6) | 76. 1 | ( 4.5 ) | 87.4 | (1.5) | . 852 | (.053) | . 978 | (.008) |
| 6-17 | 71.6 | (1.6) | 53.9 | ( 4.8) | 69.9 | (1.5) | . 753 | (.069) | . 976 | (.007) |
| 18-34 | 79.0 | (1.3) | 73.1 | ( 3.2 ) | 78.3 | (1.2) | . 925 | (.043) | . 991 | (.005) |
| 35-54 | 76.0 | (1.5) | 65.9 | ( 4.8) | 75.3 | (1.5) | . 868 | (.065) | . 991 | (.004) |
| 55-64 | 79.6 | (2.1) | 81.7 | ( 6.5) | 79.7 | (2.0) | 1.027 | (.086) | 1.001 | (.004) |
| 65 Plus | 79.9 | (1.9) | 69.5 | ( 6.7) | 79.3 | (1.8) | . 870 | (.086) | . 993 | (.005) |
| AGE OF HEAD |  |  |  |  |  |  |  |  |  |  |
| UNDER 25 | 82.5 | (2.3) | 80.9 | ( 3.6) | 82.1 | (2.0) | 381 | (.052) | . 995 | (.013) |
| $25-34$ | 82.4 | (1.3) | 69.1 | (3.7) | 80.8 | (1.2) | . 838 | (.046) | . 980 | (.006) |
| 35-44 | 75.0 | (1.5) | 54.4 | ( 4.8 ) | 73.2 | (1.4) | 725 | (.065) | . 975 | (.006) |
| $45-54$ | 75.3 | (1.6) | 70.9 | ( 5.2 ) | 75.0 | (1.5) | . 943 | (.072) | . 997 | (.004) |
| 55-64 | 78.3 | (1.9) | 74.0 | (6.2) | 78.1 | (1.8) | . 945 | (.082) | . 937 | (.005) |
| 65 PLUS | 76.1 | (1.9) | 64.2 ( | (6.2) | 75.4 | (1.8) | 844 | (.084) | . 991 | (.005) |
| SEX OF HEAD |  |  |  |  |  |  |  |  |  |  |
| MALE | 77.2 | (0.8) | 63.7 | ( 2.4) | 76.0 | (0.7) | . 825 | (.032) | .985 | (.003) |
| FEMALE | 80.2 | (1.5) | 81.6 | (3.1) | 80.4 | (1.3) | 1.019 | (.043) | 1.003 | (.006) |
| MARITAL STATUS OF HEAD |  |  |  |  |  |  |  |  |  |  |
| MARRIED | 77.5 | (0.8) | 64.8 | ( 2.6 ) | 76.5 | (0.8) | 835 | (.035) | . 987 | (.003) |
| WIDOWED | 75.1 | (2.2) | $69.2$ | ( 6.1 ) | 74.6 | (2.1) | . 922 | (.085) | . 994 | $(.007)$ |
| DIVORCED | 84.1 | (2.3) | 79.4. | ( 5.7) | 83.3 | (2.2) | . 944 | (.072) | . 990 | $(.013)$ |
| SEPARATED | 75.6 | (3.6) | $73.5$ | (5.2) | 75.1 | (3.0) | . 972 | $\text { (. } 083 \text { ) }$ | . 993 | $(.022)$ |
| NEVER MARRIED | 77.1 | (3.1) | 69.0 | (6.1) | 75.5 | (2.8) | . 895 | (.087) | . 979 | (.017) |
| FAMILY SIZE |  |  |  |  |  |  |  |  |  |  |
| ONE | 78.7 | (1.9) | 68.7 | ( 4.4) | 77.1 | (1.7) | . 874 | (.059) | . 980 | (.010) |
| TWO | 79.6 | (1.5) | $74.4$ | ( 4,5) | 79.2 | (1.4) | . 935 | (.059) | . 995 | $(.005)$ |
| THREE | 79.0 | (1.5) | $74.1$ | ( 4.1) | 78.5 | (1.4) | . 938 | (.055) | . 994 | (.006) |
| FOUR | 82.4 | (1.4) | 68.2 | ( 5.0 ) | 81.5 | (1.3) | . 827 | (.062) | . 989 | (.004) |
| FIVE | 77.7 | (1.9) | 67.5 | (6.2) | 77.0 | (1.8) | . 869 | (.082) | . 991 | $(.006)$ |
| SIX | 71.0 | (2.9) | 73.2 | (6.9) | 71.2 | (2.7) | 1.031 | (.105) | 1.002 | $(.008)$ |
| SEVEN OR MORE | 65.9 | (3.1) | 52.7 | ( 6.2) | 63.7 | (2.8) | . 801 | (. 102 ) | . 967 | (.017) |
| ADULTS IN FAMILY |  |  |  |  |  |  |  |  |  |  |
| ONE | 81.7 | (1.4) | 74.1 | ( 3.1) | 80.4 | (1.2) | 906 | (. $04 t$ ) | . 984 | (.007) |
| TWO | 78.7 | (0.9) | 66.2 | ( 2.7 ) | 77.5 | (0.9) | . 842 | (.036) | . 985 | (.003) |
| THREE | 75.0 | (1.8) | $65.3$ | ( 6.6) | 74.5 | (1.8) | 871 | (.090) | . 994 | (.005) |
| FOUR OR MORE | 72.4 | (2.6) | 60.3 | ( 9.4) | 71.7 | (2.5) | 832 | (.134) | . 989 | (.009) |
|  |  |  |  |  |  |  |  |  |  |  |
| LESS THAN \$3000 | 71.3 | (2.9) | 69.0 | ( 4.4 ) | 70.6 | (2.4) | . 968 | (.073) | . 991 | $(.021)$ |
| \$ 3000-\$4999 | 77.2 | (2.2) | 69.0 | ( 3.9) | 75.1 | (1.9) | . 894 | $(.056)$ | . 973 | $(.014)$ |
| \$ $5000-\$ 6999$ | 76.5 | (2.1) | $63.4$ | ( 4.7 ) | 74.3 | (2.0) | . 829 | $(.065)$ | . 971 | $(.011)$ |
| \$ $7000-\$ 9999$ | 75.4 | (1.9) | 61.8 | ( 4.9 ) | 73.7 | (1.8) | . 819 | (.069) | .977 .999 | (.009) |
| \$10000-\$14999 | 76.3 | (1,4) | 75.5 | ( 4.8) | 76.3 | (1.3) | . 988 | $(.066)$ | . 9995 | $(.004)$ |
| \$15000-\$24999 | 79.4 | (1.4) | 62.7 | (9.1) | 79.0 | ( 1.3 ) | . 789 | $(.116)$ | . 995 | (.003) |
| \$25000 OR MORE | 81.0 | (2.0) | 80.8 | (12.5) | 81.0 | (2.0) | 997 | (. 157 ) | 1.000 | (.002) |
| POVERTY STATUS |  |  |  |  |  |  |  |  |  |  |
| BELOW POVERTY | 72.4 | (1.9) | 66.8 | ( 2.9 ) | 70.7 | (1.6) | . 923 | (.048) | . 978 | (.014) |
| 100\% - 125\% POVERTY | 74.9 | (2.7) | 62.4 | $(6.0)$ | 73.2 | (2.5) | . 834 | $(.086)$ | . 977 | $(.012)$ |
| 125\% - 200\% POVERTY | 74.3 | (1.6) | 66.7 | ( 4.3 ) | 73.6 | (1.5) | . 898 | (.061) | .991 | $(.006)$ |
| 200\% - 300\% POVERTY | 77.2 | (1.4) | 70.5 | ( 5.1) | 76.8 | (1.4) | . 914 | (.069) | . 995 | $(.004)$ |
| 300\% - 400\% POVERTY | 79.7 | (1.6) | 73.5 | ( 7.8 ) | 79.5 | $(1.6)$ | . 922 | (.099) | . 997 | (.004) |
| $400 \%$ OR MORE POVERTY | 83.4 | (1.4) | 74.7 | ( 9.4) | 83.2 | (1.4) | . 895 | (.114) | 998 | (.003) |
| TOTAL | 77.6\% | (0.7) | $67.7 \%$ | ( 2.0) | 76.7\% | (0.6) | . 872 | (.026) | . 988 | (.002) |

"NUMEERS IN PARENTHESES ARE THE STANDARD ERROR ESTMATES.
the entire population might be a bit more disadvantaged than the data suggest. This approach would be most appropriate when working with sample sizes small enough that the bias would comprise only a small part of the total error. When a larger-scale survey is planned, the phone data might be adjusted in some way so that the entire population is approximated more closely.

I used stepwise discriminant analysis to identify which variables were most associated with differential phone coverage. The results then were used to construct a composite variable that distinguishes groups with relatively high or low phone coverage. I did this twice, once using both demographic and medical care variables and then using just demographic variables. Using the results of the discriminant analysis based on the demographic variables only, I formed the weighting categories given in Table 3. Categories like the first couple (poor persons with family heads under 25 and poor persons who are southern blacks) have weights of about 1.9 , indicating that nearly half of the persons in these groups do not have phones. Some of the last categories in the table have weights just a bit above 1.0 , indicating that nearly all persons in such groups have phones.

Tables 4 and 5 compare results when the phone population is adjusted by these weights. I also did this using the weight based on both medical care and demographic variables. The results showed the latter adjustment was not really superior to the results from adjusting by the demographic only weight, so the results are not presented here. The demographic only adjustment has the advantage of being useful to survey researchers interested in subject matters other than health.

Table 4 indicates the effect of the adjustment on basic demographic variables. The adjusted phone data approximates the data for the total population better than does the unadjusted phone data. This is the case even when the demographic variable was not used directly in the adjustment weight construction. For example, consider region, the first variable in Table 4. NonSouthSouth was the only regional distinction used in the weights to adjust the phone data. Nevertheless, the adjusted phone data is closer to the total population data for persons in the Northeast, North Central, and South; and the estimate for the Western U.S. is only slightly worse.

Table 5 and similar tables not presented here are most important, because they show the effect of the adjustment on selected health variables. Unfortunately, these tables do not show the
improvement that Table 4 showed in the distribution by demographic variables. The ratios of the total estimates to phone estimates in these tables really are not much closer to 1.0 than are the ratios using the unadjusted phone data. In Table 5, presenting data on the percent contacting a doctor during the year, the ratio for the total population is virtually unchanged. None of the subgroups shows any real improvement with the adjustment.

There are a few estimates in tables not shown here that are improved by the adjustment, especially in the estimated percent who were completely satisfied with their most recent medical visit. For example, the ratio for the total population improves to .995 , while the unadjusted ratio was .985. However, in general I would say that the adjustment process allows the adjusted phone data to approximate the total data fairly well in terms of demographics, but it still provides estimates that somewhat overrepresent the health-advantaged population.

I suppose that these results are not that unexpected, for two reasons. One reason is that, as was said earlier, the correlation between phone coverage and the amount of difference in medical care estimates is not as great as generally supposed. However I had hoped that the correlation would be large enough that an adjustment by coverage rates would make a substantial improvement in the phone population estimates.

The second reason that these results are not all that unexpected is that the ratios were fairly close to 1.0 in the unadjusted data, even though many were significantly different from 1.0 statistically. Because they were so close to 1.0 , there really was not much room for improvement.

When running the stepwise discriminant analysis which included health care variables, the only medical care variable which had large Fs was a three-category insurance variable: without insurance, with Medicaid or other reduced price insurance, and with regular group or individual insurance coverage. The Fs of the dependent variables in Table 5 and other tables not shown here (contact with a doctor and so on) were quite low. This indicates that other differences between the phone and nonphone populations were more important.

However, if possible. I would like suggestions on any other avenues to explore in terms of adjusting the phone population data so that it better approximates the total population.

| CATEGORY | POVERTY STATUS | $\begin{gathered} \text { AGE OF } \\ \text { HEAD } \end{gathered}$ | RACE | REGION | ADULTS IN FAMILY, MARITAL STATUS OF HEAD | RESIDENCE | $\begin{aligned} & \text { PERCENT } \\ & \text { OF TOTAL } \\ & \text { POPULATION } \end{aligned}$ | WEIGHT |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 1 | poor | under 25 | al1 | all | al1 | al 1 | 1. $2 \%$ | 1.8810 |
| 2 | poor | 25 plus | So. Black | all | a 11 | al1 | 1.3 | 1.9472 |
| 3 | poor | 25 plus | Spanish SW | all | al 1 | al 1 | 1.6 | 1.5998 |
| 4 | poor | 25-34 | other | South | al 1 | all | 0.9 | +.9816 |
| 5 | poor | 35-64 | other | South s | al 1 | all | 3.1 | 1.4039 |
| 6 | poor | 65 plus | other | South | all | all | 1.0 | f. 1532 |
| 7 | poor | 25-24 | other | nonsouth | a11 | al 1 | 1.3 | 1.4460 |
| 8 | poor | $35-64$ | other | nonsouth | 2 plus, not sep. or div. | all | 2.1 | 1.1550 |
| 9 | poor | 35-64 | other | nonsouth | all | al1 | 0.8 | 1.2987 |
| 10 | poor | 65 plus | other | nonsouth | all | al1 | 1.3 | 1.0456 |
| 11 | nonpoor | under 25 | a 11 | South | all | all | 1.7 | 1.5484 |
| 12 | nonpoor | under 25 | all | nonsouth | 2 plus, not sep. or div. | a. 11 | 2.6 | +. 1584 |
| 13 | nonpoor | under 25 | all | nonsouth | other than above | al 1 | 0.9 | 1. 1793 |
| 14 | nonpoor | 25-34 | al1 | South | 2 plus, not sep. or div. | a 11 | 4.8 | 1.1836 |
| 15 | nompoor | 25-34 | al 1 | South | other than above | al 1 | 0.9 | 1.4978 |
| 16 | nompoor | 25-34 | all | nonsouth | separated or divorced | al1 | 1.3 | 1.0748 |
| 17 | nompoor | 25-34 | al 1 | nonsouth | 1. not separated or div. | al1 | 0.9 | 1.0706 |
| 18 | nonpoor | 25-34 | al1 | nonsouth | 2 , not separated or div. | rural nonfarm | 1.6 | 1.0330 |
| 19 | nonpoor | 25-34 | al1 | nonsouth | 2 , not separated or div. | other | 8.8 | 1.0326 |
| 20 | nonpoor | 25-34 | all | nonsouth | 3 plus, not sep. or div. | al 1 | 0.8 | 1.0651 |
| 21 | nonpoor | 35 pius | So. Black | a 11 | a 11 | all | 0.8 | 1.2073 |
| 22 | nonpoor | 35 plus | Spanish Sw | al 1 | all | al 1 | 1.4 | 1.1337 |
| 23 | nompoor | 35 plus | other | South | separated or divorced | all | 0.9 | 1.2004 |
| 24 | nompoor | 35-64 | other | South | 1. not separated or div. | all | 0.6 | 1.1008 |
| 25 | nonpoor | $35-64$ | other | South | 2 , not separated or div. | all | 7.0 | 1.0425 |
| 26 | nompoor | 35-64 | other | South | 3 , not separated or div. | all | 4.2 | 1.0376 |
| 27 | nonpoor | 35-64 | other | South | 4 plus, not sep. or div. | all | 1.7 | 1.1283 |
| 28 | nonpoor | 65 plus | other | South | a11 | al 1 | 2.1 | 1.0737 |
| 29 | nonpoor | $35-64$ | other | nonsouth | 1, separated or div. | all | 1.2 | 1.0886 |
| 30 | nonpoor | 35-64 | other | nonsouth | 2 plus, sep. or div. | all | 1.4 | 1.0556 |
| 31 | nonpoor | 35-64 | other | nonsouth | 1, not separated or div. | all | 1.7 | 1.0700 |
| 32 | nonpoor | 35-64 | other | nonsouth | 2, not sep. or div. | rural nonfarm | 2.8 | 1.0301 |
| 33 | 1-2 poverty | 35-64 | other | nonsouth | 2 , not sep. or div. | other | 2.8 | 1.0414 |
| 34 | $2+$ poverty | 35-64 | other | nonsouth | 2, not sep. or div. | other | 10.9 | 1.0231 |
| 35 | nompoor | $35-64$ | other | nonsouth | 3 , not sep. or div. | rural nonfarm | 1.5 | 1.0000 |
| 36 | nonpoor | 35-64 | other | nonsouth | 3, not sep. or div. | other | 7.5 | 1.0339 |
| 37 | nonpoor | 35-64 | other | nonsouth | 4 plus, not sep. or div. | al1 | 6.5 | 1.0143 |
| 38 | nonpoor | 65 plus | other | nonsouth | 2 plus, not sep. or div. | al 1 | 4.9 | 1.0188 |
| 39 | nonpoor | 65 plus | other | nonsouth | other than above | all | 1.8\% | 1.0506 |

TABLE 4 - PERCENT OF THE POPULATION BY DEMOGRAPHIC CHARACTERISTICS, TOTAL POPULATION AND UNADUUSTED AND ADJUSTED PHONE POPULATIONS; CHAS 1976

| PHONE POPULATION |  |  |  |  |  | PHONE POPULATION |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| CHARACTERISTIC | TOTAL POPULATION | UNADJUSTED | ADJUSTED | CHARACTERISTIC | TOTAL POPULATION | UNADJUSTED | ADJUSTED |
| REGION |  |  |  | MARITAL STATUS OF H | HEAD |  |  |
| NORTHEAST | 22.39\% | 23.46\% | 22.47\% | MARRIED | 80.85 | 82.30 | 81.04 |
| NORTH CENTRAL | 30.57 | 32.23 | 30.83 | WIDOWED | 6.77 | 6.87 | 6.97 |
| SOUTH | 32.67 | 30.03 | 32.61 | DIVORCED | 5.41 | 4.94 | 5.36 |
| WEST | 14.37 | 14.28 | 14. 10 | SEPARATED | 3.50 | 2.85 | 3.20 |
| RESIDENCE |  |  |  | NEVER MARRIED | 3.46 | 3.06 | 3.43 |
| SMSA CENTRAL CITY | 25.59 | 25.47 | 25.95 | FAMILY SIZE |  |  |  |
| SMSA OTHER | 36.65 | 37.75 | 36.90 | ONE | 7.56 | 6.99 | 7.31 |
| NONSMSA URBAN | 11.74 | 11.49 | 11.63 | TWO | 19.54 | 19.86 | 19.45 |
| RURAL NONFARM | 20.14 | 19.18 | 19.41 | THREE | 17.03 | 16.89 | 16.90 |
| RURAL FARM | 5.88 | 6. 11 | 6.11 | FOUR | 20.31 | 20.98 | 20.94 |
| RACE |  |  |  | FIVE | 15.96 | 16.40 | 16.27 |
| SPANISH HERITAGE, |  |  |  | SIX | 8.88 | 9.06 | 8.89 |
| SOUTHWEST | 4. 14 | 3.37 | 3.96 | SEVEN OR MORE | 10.71 | 9.82 | 10.28 |
| OTHER WHITE | 83.82 | 85.86 | 84.21 | ADULTS IN FAMILY |  |  |  |
| NONSMSA SOUTHERN |  |  |  | ONE | 13.88 | 12.67 | 13.59 |
| BLACK | 2.44 | 1.62 | 2.29 | TWO | 55.95 | 55.89 | 55.47 |
| OTHER NONWHITE | 9.60 | 9.13 | 9.54 | THREE | 18.56 | 19.44 | 18.93 |
| AGE |  |  |  | FOUR OR MORE | 11.63 | 12.00 | 12.01 |
| 0-5 | 9.32 | 8.78 | 9.29 | FAMILY INCOME |  |  |  |
| $6-17$ | 24.59 | 24.48 | 24.44 | LESS THAN \$3000 | 5.29 | 4.20 | 5.04 |
| 18-34 | 24.97 | 24. 17 | 24.95 | \$ 3000-\$4999 | 8.93 | 7.37 | 8.43 |
| 35-54 | 21.76 | 22.34 | 21.68 | \$ 5000-\$6999 | 9.93 | 8.61 | 9.27 |
| 55-64 | 9.35 | 9.80 | 9.55 | \$ $7000-\$ 9999$ | 12.61 | 12.13 | 12.38 |
| 65 PLUS | 10.03 | 10.43 | 10.09 | \$10000-\$14999 | 25.47 | 26.38 | 25.71 |
| AGE OF HEAD |  |  |  | \$15000 - \$24999 | 25.94 | 27.90 | 26.54 |
| UNDER 25 | 6.35 | 5. 19 | 6.35 | \$25000 OR MORE | 12.36 | 13.42 | 12.63 |
| 25-34 | 21.67 | 21.02 | 21.66 | PQVERTY STATUS |  |  |  |
| 35-44 | 26.46 | 26.55 | 26.02 | BELOW POVERTY | 14.51 | 11.41 | 14.51 |
| 45-54 | 20.99 | 21.76 | 21.12 | 100\%-125\% PQVERTY | 6.69 | 6.38 | 6.26 |
| 55-64 | 12.72 | 13.22 | 12.91 | 125\%-200\% POVERTY | 20.08 | 20.10 | 19.57 |
| 65 PLUS | 11.81 | 12.27 | 11.94 | 200\%-300\% POVERTY | 23.98 | 24.92 | 24.03 |
| SEX OF HEAD |  |  |  | 300\%-400\% POVERTY | 16.62 | 17.66 | 17.05 |
| MALE | 85.16 | 85.92 | 84.82 | 400\% OR MORE |  |  |  |
| FEMALE | 14.84 | 14.08 | 15.18 | POVERTY | 18. $12 \%$ | 19.52\% | 18.58\% |

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| REGION |  |  |  |  |  | MARITAL STATUS OF H | HEAD |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| NORTHEAST | 81.4\% | 81.4\% (1.4) | . 996 | . 996 | (.004) | MARRIED | $77.5 \%$ | 77.5\%(0.8) | . 987 | 987 | (.003) |
| NORTH CENTRAL | 76.8 | 76.6 (1.3) | 1.001 | 1.004 | (.003) | WIDOWED | 74.3 | 75.1 (2.2) | 1.003 | 994 | (.007) |
| SOUTH | 75.0 | 74.8 (1.2) | .979 | . 981 | (.006) | DI VORCED | 84.1 | 84.1 (2.3) | . 990 | . 990 | (.013) |
| WEST | 78.8 | 79.2 (1.5) | . 972 | 966 | (.006) | SEPARATED | 74.9 | 75.6 (3.6) | 1.002 | 993 | (.022) |
| RESIDENCE |  |  |  |  |  | NEVER MARRIED | 77.0 | 77.1 (3.1) | . 980 | 979 | (.017) |
| SMSA CENTRAL CITY | 78.1 | 78.1 (1.3) | . 989 | 989 | (.005) | FAMILY SIZE |  |  |  |  |  |
| SMSA OTHER | 79.9 | 79.9 (1.1) | . 990 | 990 | (.004) | ONE | 78.6 | 73.7 (1.9) | . 980 | . 980 | (.040) |
| NONSMSA URBAN | 74.6 | 75.1 (1.9) | . 997 | 990 | (. 0007 ) | TWO | 79.6 | 79.6 (1.5) | . 994 | . 995 | (.005) |
| RURAL NONFARM | 76.1 | 76.0 (1.6) | . 982 | . 983 | (.007) | THREE | 79.2 | 79.0 (1.5) | . 992 | . 994 | (.006) |
| RURAL FARM | 71.2 | 70.6 (3.2) | . 986 | 994 | (.009) | FOUR | 82.3 | 82.4 (1.4) | . 990 | . 989 | (.004) |
| RACE |  |  |  |  |  | FIVE | 77.4 | 77.7 (1.9) | . 995 | . 991 | (.006) |
| SPANISH HERITAGE, |  |  |  |  |  | SIX | 70.5 | 71.0 (2.9) | 1.009 | 1.002 | (.008) |
| SOUTHWEST | 71.2 | 72.0 (5.7) | . 908 | . 898 | (. 042 ) | SEVEN OR MORE | 66.7 | 65.9 (3.1) | 954 | 967 | (.017) |
| OTHER WHITE | 78.1 | 78.0 (0.8) | . 993 | 995 | (.003) | ADULTS IN FAMILY |  |  |  |  |  |
| NONSMSA SOUTHERN |  |  |  |  |  | ONE | 81.4 | 81.7 (1.4) | 987 | 984 | (.007) |
| BLACK | 69.0 | 69.6 (4.7) | . 943 | 936 | (.043) | TWO | 78.6 | 78.7 (0.9) | . 986 | 985 | (.003) |
| OTHER NONWHITE | 77.5 | 77.5 (3.8) | . 994 | . 994 | (.016) | THREE | 75.0 | 75.0 (1.8) | . 994 | . 994 | (.005) |
| AGE |  |  |  |  |  | FOUR OR MORE | 72.4 | 72.4 (2.6) | . 990 | 989 | (.009) |
| $0-5$ | 89.7 | 89.3 (1.6) | . 974 | 978 | (.008) | FAMILY INCOME |  |  |  |  |  |
| 6-17 | 71.5 | 71.6 (1.6) | . 978 | . 976 | (.007) | LESS THAN \$3000 | 71.9 | 71.3 (2.9) | . 982 | . 991 | (.021) |
| 18-34 | 79.0 | 79.0 (1.3) | . 991 | . 991 | (.005) | \$3000-\$4999 | 77.3 | 77.2 (2.2) | . 972 | . 973 | (.014) |
| $35-54$ | 76.1 | 76.0 (1.5) | . 990 | . 991 | (.004) | \$ 5000-\$6999 | 76.4 | 76.5 (2.1) | . 972 | . 971 | (.011) |
| 55-64 | 79.5 | 79.6 (2.1) | 1.003 | 1.001 | (.004) | \$ $7000-\$ 9999$ | 75.3 | 75.4 (1.9) | . 979 | . 977 | (.009) |
| 65 PLUS | 79.8 | 79.9 (1.9) | . 994 | . 993 | (.005) | \$10000-\$14999 | 76.4 | 76.3 (1.4) | . 999 | . 999 | (.004) |
| AGE OF HEAD |  |  |  |  |  | \$15000-\$24999 | 79.6 | 79.4 (1.4) | . 993 | . 995 | (.003) |
| UNDER 25 | 82.0 | 82.5 (2.3) | 1.001 | 995 | (.013) | \$25000 OR MORE | 81.0 | 81.0(2.0) | . 999 | 1.000 | (.002) |
| 25-34 | 82.3 | 82.4 (1.3) | . 982 | 980 | (.006) | POVERTY STATUS |  |  |  |  |  |
| 35-44 | 74.9 | 75.0 (1.5) | . 977 | . 975 | (.006) | BELOW POVERTY | 72.9 | 72.4 (1.9) | . 971 | . 978 | (.014) |
| 45-54 | 75.1 | 75.3 (1.6) | . 998 | . 997 | (.004) | 100\%-125\% POVERTY | Y 75.0 | 74.9 (2.7) | . 975 | . 977 | (.012) |
| 55-64 | 78.2 | 78.3 (1.9) | . 998 | . 997 | (.005) | 125\%-200\% POVERTY | Y 74.4 | 74.3 (1.6) | . 989 | . 991 | (.006) |
| 65 PLUS | 75.9 | 76.1 (1.9) | . 994 | . 991 | (.005) | 200\%-300\% POVERTY | Y 77.4 | 77.2 (1.4) | . 992 | . 995 | (.004) |
| SEX OF HEAD |  |  |  |  |  | 300\%-400\% POVERTY | Y 79.7 | 79.7 (1.6) | . 997 | . 997 | (.004) |
| MALE | 77.2 | 77.2 (0.8) | . 985 | . 985 | $(.003)$ | 400\% OR MORE |  |  |  |  |  |
| FEMALE | 79.9 | 80.2 (1.5) | 1.006 | 1.003 | (.006) | POVERTY | 83.5 | 83.4 (1.4) | .997 | . 998 | (.003) |
|  |  |  |  |  |  | TOTAL | 77.5\% | $77.6 \%(0.7)$ | . 989 | . 988 | (.002) |

*NUMBERS IN PARENTHSES ARE THE STANDARD ERROR ESTMATES.


[^0]:    *NUMBERS IN PARENTHESES ARE THE STANDARD ERROR ESTIMATES FOR BOTH THE PHONE AND THE NONPHONE POPULATIONS.

