INTRODUCTION: Controlled selection is a probability sampling procedure which enables its users to purposively introduce constraints on the distribution of the sample. There is also an expectation of reduced sample variances relative to those from other stratified sample designs.

For more than 30 years, the University of Michigan's Survey Research Center has used this technique in developing samples for research studies. The Bureau of the Census has used controlled selection for the Current Population Survey samples in order to control sample distributions by states (U.S. Bureau of the Census, 1963, 1978). A past revision of the Consumer Price Index sample also applied controlled selection techniques to achieve rigorous geographic control (Wilkerson,1961). Gains in precision resulting from controlled selection have been explored, but not enough is yet known about that aspect of potential benefits from the use of controlled selection.

The purpose of a current research investigation at the Survey Research Center is twofold: (1) To observe the performance of various forms of controlled selection when compared among themselves as well as with other selection modes; (2) To investigate ways to improve the application of controlled selection to sampling methods.

The purpose of this paper is to give a progress report on some of the research undertaken to date.

There are two principal findings: (1) For the same sample design, computer generated controlled selection often leads to slightly higher variances than does manual controlled selection; but since the differences in precision are small and manual controlled selection is laborious, computer generated controlled selection is preferred; (2) With stratified random sampling as abasis for comparison, computer generated controlled selection, and ordered systematic selection both result in lower between primary sampling unit variances, and the larger reduction generally is effected by controlled selection.
DATA USED IN THE RESEARCH: Research populations: In order to simulate sampling operations, data from the 1960 census were employed in developing alternative sample designs that were then tested with census data from later periods. Counties or county groups in the North Central and South Regions of the United States were regarded as separate populations.

Excluded from each region are the Standard Metropolitan Statistical Areas (SMSA's) with 1960 populations in excess of 500,000 inhabitants. It is assumed that those areas would be included with certainty in any regional sample. The proportion of 1960 inhabitants remaining in each region after excluding major urban areas was a little over half in the North Central and about three-fourths in the South.
Farmation of primary sampling units: The primary sampling units are SMSA's, single counties or groups of geographically contiguous counties combined to meet a minimum size of approximately 20,000 inhabitants in 1960. In the North Central 777 psu's were formed, in the South 1,046 .
Stratification: In addition to the separation of SMSA's and non-SMSA's, two other dimensions in stratification were observed: geographic location
and level of urbanization. Geographic classifications were states or state groups. Urban categories were developed from the 1960 census reports of percent urban for counties and SMSA's. Selection probabilities: In order to proceed with the research, it was necessary to specify sample sizes. In the North Central 40 selections are assumed, with exactly 12 from SMSA's and 28 from non-SMSA's. In the South 56 selections are assumed, 18 SMSA's and 38 non-SMSA's. The sizes remain constant for all investigations. Selection probabilities were calculated separately for SMSA's and non-SMSA's of each region.
The choice of variables: The aggregates listed in Table 1 were selected from the 1972 County City Data Book [U.S. Bureau of the Census, 1973], with the intent to include varibles describing persons, housing units or households, farms, businesses and industries. Census data reported as percentages were converted to aggregates and regarded as exact values thereafter.

Table 2 contains ratios of some of the aggregates listed in Table 1.
CONTROLLED SELECTION DESIGNS USED IN THE RESEARCH: Different forms of stratification for controlled selection as well as a simple example of the procedure are given in Appendix A. Three controlled selection designs, all of one type, have been used so far in the research. For the type used, selection probabilities sum to an integer over rows and columns but not necessarily by rows or by columns. Designs I and II relate to the North Central Region, III to the South.

Design I is a cross-tabulation of six state groups by 11 urban classes - 7 for non-SMSA's and 4 for SMSA's. The psu probabilities sum to exactly 28 over all non-SMSA's, 12 over all SMSA's. In the matrix of 66 cells, 49 are nonzero.

Design II differs from I by increasing the state categories from six to 12 , and the matrix from 66 to 132 cells, of which 86 are nonzero. Design II generally has lower variances--a result of increased stratification.

Design III in the South is similar to the North Central designs. The psu's were assigned to seven state groups and 12 urban categories. There are 69 nonzero cells.
MANUAL CONTROLLED SELECTION COMPARED WITH COMPUTER CONTROLLED SELECTION: Does the application of controlled selection to sampling procedures require the services of an experienced sampler, or can computers be programmed to perform the operations satisfactorily? The displays in Tables 3 and 4 are responses to that question. Both manual and computer controlled selections were completed for each of the three designs. Performance tests are based on the precision of estimates achieved by each process.

The relative variances from manual controlled selections are the bases with which relative variances from computer generated controlled selections are compared. Variable code numbers in Table 3 correspond to those in Table 1 where each variable is described. Similarly, the ratios fm Table 4 are described in Table 2.

The variances reported in Tables 3 and 4 are between psu relative variances which are population values. No sampling of primary units occurs.

In columns 3, 5 and 7 of Table 3, a quotient in excess of one indicates higher variance and lower precision from computer selections. The average quotients and ranges are given at the bottom of the table. Neither procedure shows a clear advantage.

| Table 1. AGGREGATES IHCLUDED IN NUMERICAL CALCULATIONS NORTH CENTRAL AND SOUTH REGIONS OF THE UNITED STATES EXCLUSIVE OF MAJOR METROPOLITAN AREAS |  |  |  |
| :---: | :---: | :---: | :---: |
| Vari <br> able <br> Code <br> Nos. | Variable Descriptions | Regions* |  |
|  |  | North Central (part) | $\begin{aligned} & \text { South } \\ & \text { (part) } \end{aligned}$ |
| 1 | Population, 1970 | 31,466,373 |  |
| 2 | Civilian labor force | 12,232,279 | 16,616,911 |
| 3 | Females | 4,460,926 | 6,353,671 |
| 4 | Married, husband present | 2,713,839 | 3,951,887 |
| 5 | Persons 65 years and ove | 3,505,126 | 4,559,033 |
| 6 | Persons below low income level in 1969 | 3,695,015 | 10,091,864 |
| 7 | Persons 3-34 years enrolled in college | 1,177,522 | 1,223,663 |
| 8 | Farm population | 3,853,953 | 2,899,976 |
| 9 | Persons below low. income level in 1969 | 572,608 |  |
| 10 | Black population | 841,593 | 8,316,842 |
| 11 | Persons of Spanish heritage | 229,368 | 1,553,119 |
| 12 | Vote cast for President, 1968 | 12,451,970 | 14,137,207 |
| 13 | Families, 1970 | 7,911,202 | 11,508,220 |
| 14 | With income $\$ 25,000$ or more in 1969 | 246,806 | 282,872 |
| 15 | Recipients of aid to families with dependent children, Feb. 1972 | 994,005 | 2,012,089 |
| 16 | Public assistance payments, Feb. 1972 ( $\$ 1,000$ ) | 82,015 | 130,899 |
| 17 | Year-round housing units, 1970 | 10,506,469 | 15,065,608 |
| 18 | Occupied housing units | 9,730,068 | 13,760,894 |
| 19 | With telephone availabl | 8,798,148 | 10,771,490 |
| 20 | With home food freezer | 3,974,586 | 4,895,629 |
| 21 | Household head moved into unit during 1965-70 | 4, 464,355 | 6,853,436 |
| 2 | One person households | 1,677,300 | 2,109,300 |
| 23 | Farms, 1969 Census of Agriculture |  | 1,129,186 |
| 24 | With sales \$2,500 and over | 844,926 | 566,972 |
| 25 | Value of farm products sold by farms with sales of $\$ 2,500$ and over, 1969 ( $\$ 1,000$ ) | 19,238,295 | 12,387,695 |
| 26 | Value of livestock and livestock products sold, 1969 (\$1,000) | 10,084,211 | 4,041,223 |
| 27 | Manufacturing estab's, 1967 Census of Manufactures | 42,390 | 56,826 |
| 28 | Establishments with 100 or more employees | 5,263 | 7,395 |
| 29 | Mineral industries estabIishments, 1967 Census of Mineral Industries** | 8,612 | 24,020 |
| 30 | Retail trade establishments 1967 Census of Business | 314,083 | 427,658 |
| 31 | Sales of estab's ( $\$ 1,000$ ) | 48,863,904 | 58,558,573 |
| 32 | Estab's with payroll | 221,488 | 271,097 |
| * Entries in these columns will not agree with figures shown in Census Bureau publications, for three reasons: 1) major metropolitan areas have been excluded; 2) to avoid dividing SMSA's that crossed regional boundaries, a clear separation of regions was sacrificed; 3) where data of interest to the research were reported as percents in the County and City Data Book, they were converted to aggregates and regarded thereafter as exact values. ** Mineral industries establishments shown in this table are sums of county totals, which do not agree with state totals reported In the County and City Data Book. Source: U.S. Bureau of the Census, 1973 . |  |  |  |

Turning to Table 4, consider the performance of each selection design when estimates are ratios. Averages of quotients, given below the table, are a little more than one, indicating some increased variance on the average from computer generated controlled selection.

Over all observations, manual controlled selection appears to give slightly higher precision within the bounds of the tests. Nevertheless, the authors agree that the laborious process of manual controlled selection has hindered its

Table 2. RATIOS INCLUDED IN NUMERICAL CALCULATIONS, NORTH CENTRAL AND SOUTH REGIOHS OF THE UNITED STATES EXCLUSIVE OF MAJOR METROPOLITAN AREAS

| Ratio Code Nos. | Ratio Descriptions | Num. $\varepsilon$ Denom Code Nos. ${ }^{*}$ | Regions \%\% |  |
| :---: | :---: | :---: | :---: | :---: |
|  |  |  | $\begin{gathered} \text { North } \\ \text { Central } \\ \text { (part) } \\ \hline \end{gathered}$ | $\begin{array}{\|l\|l} \text { South } \\ \text { (part) } \end{array}$ |
| 40 | Percent of total population voting for President, 1968 |  | 39.572 | 31.223 |
| 41 | Percent Black poplation |  | 5 |  |
| 42 | Percent of persons 65 years of age and over | ,1 | 11. |  |
| 43 | Percent Spanish heritage population | 11,1 | 0.729 | 3.430 |
| 44 | Percent persons 3-34 years enrolled in college | 7,1 | 3.742 | 2.703 |
| 45 | Percent females in civilian labor force | 3,2 | 36.468 | 38.236 |
| 46 | Percent females in labor force married with husband present | 4,3 | 60.836 |  |
| 47 | Percent families with income $\$ 25,000$ or more | 14,13 | 3.120 |  |
| 48 | Average public assistance payment per family | 16,13 | \$10.367 | \$11.374 |
| 49 | Percent farm poputation below low income level | 9,8 | 14.858 | . 042 |
| 50 | Percent occupied housing units | 18,17 |  |  |
| 51 | Percent occupancy by one-person households | 22,18 | 17.238 | 15.328 |
| 52 | Percent occupancy by movers into units during 1965-1970 | 21,18 | 45.882 | 804 |
| 53 | Percent occupancy with home freezer |  | 40.848 |  |
| 54 | Percent occupancy with telephone available | 19,18 | 90.422 |  |
| 55 | Percent farms with |  |  |  |
| 56 | sales $\$ 2,500$ or more Livestock and livestock products sales as percent of total sales by farms with sales of \$2,500 or more | 24,23 | 75.880 52.417 | 50.211 32.623 |
| 57 | Percent manufacturing establishments with 100 or more employees | 28,27 | 12.416 | . 013 |
| 58 | Average sales per retail trade establishment ( $\$ 1,000$ ) | 31,30 | 155.576 | 29 |
| 59 | Percent retail trade establishments with payroll | 32,30 | 70.519 | 63.391 |

acceptance and probably would continue to do so, whereas controlled selection by computers gives acceptable precision and places the sampling technique on a practical basis.
CONTROLLED SELECTION COMPARED WITH STRATIFIED RANDOM SAMPLING AND ORDERED SYSTEMATIC SAMPLING: How does the precision of estimates from controlled selection compare with that from other selection methods? Comparisons are made with two other sampling procedures: stratified random sampling, and ordered systematic sampling, each method deTable 3. BETUEEN PSU RELATIVE VARIANCES OF AGGREGATES FOR three manuai controlled selection designs, and comparisons WITH RELATIVE VARIANCES FOR COMPUTER GENERATED CONTROLLED SELECTIONS

| 1- |  | North | Central |  |  | , |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| able | Des | n It | Des | ign IIf | Desigi | IIIf |
| de |  | $\mathrm{v}^{2}{ }^{\text {d }}{ }^{5}$ |  |  |  |  |
|  |  | $\frac{\text { cs-comip }}{\text { V2 }}$ |  | $\frac{y^{2}}{}$ | $\mathrm{v}^{2} \ddagger$ | $y_{\text {cocomp }}$ |
| Nos.* | yct-man | $\overline{v_{u}^{2}}{ }^{2} \ddagger \text { man }$ | CSTman | $\overline{v_{c}} \bar{y}_{\substack{2}}$ | $y_{\text {cs-mar }}$ | $\frac{v_{c}^{2}+\frac{c o m p}{\ddagger}}{y_{\text {man }}}$ |
| 1 | 2 | 3 | 4 | 5 | 6 | 7 |
| 1 | . 000350 | 1.003 | . 000359 | . 966 | . 000525 | 1.009 |
| 2 | . 000519 | 1.002 | . 000559 | . 887 | . 000754 | . 987 |
| 3 | . 000777 | . 991 | . 000851 | . 861 | . 000866 | . 981 |
| 4 | . 000760 | . 992 | .000813 | . 895 | . 001109 | . 992 |
| 5 | . 000822 | . 995 | . 000748 | 1.052 | . 002789 | 1.028 |
| 6 | . 002309 | 1.008 | . 002017 | . 986 | . 001414 | . 996 |
| 7 | . 041173 | . 985 | . 041290 | . 951 | . 028098 | 1.005 |
| 8 | . 004744 | 1.031 | . 004208 | 1.007 | . 010901 | 1.004 |
| 9 | 014536 | 1.015 | . 013022 | . 994 | . 022607 | . 999 |
| 10 | 020330 | 1.091 | . 018800 | 1.109 | . 006271 | . 989 |
| 11 | . 038245 | 1.064 | . 037380 | 1.031 | . 123861 | 1.006 |
| 12 | . 000353 | . 986 | . 000360 | . 983 | . 000849 | 1.003 |
| 13 | 000305 | . 999 | . 000317 | . 977 | . 000649 | 1.013 |
| 14 | 003669 | 1.006 | . 003607 | . 896 | . 005025 | 1.012 |
| 15 | 017069 | 1.013 | -015270 | 1.009 | . 003780 | . 996 |
| 16 | 012085 | 1.000 | . 009511 | 1.002 | . 003144 | 1.034 |
| 17 | 000439 | . 992 | . 000422 | 1.033 | . 000714 | 1.027 |
| 18 | . 000288 | . 997 | . 000302 | . 960 | . 000658 | 1.020 |
| 19 | . 000341 | 1.011 | 000355 | . 966 | . 000928 | 1.027 |
| 20 | . 000799 | 1.007 | 000793 | . 986 | . 001120 | 1.016 |
| 21 | . 000922 | 1.014 | . 000905 | . 950 | . 001486 | 1.012 |
| 22 | . 000621 | . 934 | 000560 | 1.031 | . 001238 | 1.048 |
| 23 | . 004278 | 1.021 | . 004109 | . 973 | . 007984 | 1.009 |
| 24 | . 005736 | 1.034 | 004944 | 1.022 | . 012578 | 1.026 |
| 25 | . 013664 | 1.028 | . 010298 | 1.023 | . 046291 | 1.012 |
| 26 | . 029450 | 1.009 | . 022123 | 1.006 | . 206458 | 1.005 |
| 27 | 003697 | . 970 | . 003362 | . 971 | . 004404 | 1.013 |
| 28 | 007021 | . 986 | . 006523 | . 990 | . 007580 | 1.002 |
| 29 | . 169167 | . 995 | . 148758 | 1.002 | . 071519 | 1.002 |
| 30 | - 000779 | 1.031 | . 000758 | . 981 | . 000662 | 1.005 |
| 31 | 000710 | 1.057 | 000653 | 1.055 | . 000929 | . 956 |
| 32 | 000747 | 1.015 | , 000721 | . 991 | . 000751 | 1.009 |
|  |  | Sum | mary Meas | sures |  |  |
| rithr | tic means | 1.009 | NA | . 986 | NA | 1.008 |
| Range |  | $\begin{aligned} & .970-091 \end{aligned}$ | NA | $\begin{aligned} & .861- \\ & 1.109 \end{aligned}$ | NA | $.956-$ |

WA Not applicable. * See Table 1 for variable descriptions. + See Text for design descriptions. Design I has 40 selec tions from 49 cells; Design II has 40 selections from 86 cells; Design III has 56 selections from 69 cells.

signed to have exactly two selections per stratum. These two methods differ only in the way primary units are paired within strata. The procedures are discussed and illustrated in Appendix B.

The stratified designs use the same primary units, the same stratification variables, the same number of sample selections, and with few exceptions the same psu probabilities as were used for the controlled selection designs. Occasionally, psu probabilities were adjusted slightly so that their sums within strata would be exactly two. No psu crosses a stratum boundary. Thestratified designs share the same strata, which were constructed to satisfy the ordered systematic design.

For the randomized design, psu's within strata were rearranged in a random order as described in Appendix $B$.

The ordered systematic design may be regarded as a form of controlled selection, or, as Goodman and Kish discussed in their 1950 paper, controlled selection may be viewed as systematic sampling when the primary units are ordered in meaningful sequence.

| $\begin{array}{l\|} \hline \text { Varli } \\ \text { able } \\ \text { code } \\ \text { Nos. } \end{array}$ | North Central |  |  |  | South |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Design It |  | ${ }^{\frac{1}{2}}$ |  | Design IIIf |  |
|  | $v_{y_{c-m a n}^{2}}^{\ddagger}$ |  |  |  | $v_{y_{c s-m a r}}^{2} \ddagger$ |  |
|  |  | $\frac{y \text { cs-comp }}{v^{2}+}$ |  | $\frac{\text { cs-comp }}{\mathrm{v}^{2}}$ |  |  |
|  |  | $V^{v_{c}^{2}} \pm \pm \text { man }$ | chman | $\bar{v}^{2} y_{c}^{ \pm} \pm \text {man }$ |  |  |
| 1 | 2 | 3 | 4 | 5 | 6 |  |
| 40 | . 000300 | . 990 | . 000267 | . 037 | . 0005 | . 014 |
| 41 | . 020502 | 1.097 | . 0189 | 1.112 | . 0069 | . 007 |
| 42 | . 001101 | 982 | . 0009 | . 03 | . 0023 | . 016 |
| 43 | . 037387 | 1.064 | . 036438 | 1.042 | . 123578 | 1.006 |
| 44 | . 038384 | . 984 | . 038454 | . 956 | . 026443 | 1.001 |
| 45 | . 000099 | . 970 | . 00010 | . 888 | . 000090 | 1.051 |
| 46 | . 000094 | 1.000 | . 00009 | 981 | . 0000 | 1.034 |
| 47 | . 003011 | 1.008 | . 00286 | . 921 | . 003305 | 1.015 |
| 48 | . 012347 | 1.000 | . 00976 | . 004 | . 00407 | 1.028 |
| 49 | . 007202 | 1.018 | . 0064 | 999 | . 0074 | 1.008 |
| 50 | . 000096 | 1.001 | . 0000 | 1.038 | . 00002 | . 018 |
| 51 | . 000411 | 927 | . 0003 | 1.029 | . 0004 | . 060 |
| 52 | . 000387 | 1.028 | . 00036 | 988 | . 000 | . 008 |
| 53 | . 000551 | . 988 | . 0005 | 932 | . 0009 | . 060 |
| 54 | . 000040 | 1.030 | . 00003 | . 09 | 000 | . 2.26 |
| 55 | . 000886 | 1.01 | . 0005 | 1.11 | . 00 | 02 |
| 56 | 006269 | . 996 | . 00508 | 99 | 800 | 1.001 |
| 57 | 004157 | 998 | . 0038 | 1.09 |  | . 98 |
| 58 | $\begin{array}{r} .000742 \\ .000106 \end{array}$ | . 02 | . 000 | . 956 | . 000533 | . 961 |
| 59 |  | 1.031 | . 000084 | 1.059 | . 000108 | 1.019 |
| Summarv |  |  |  |  |  |  |
| Arithmetic means Ranges |  | $\begin{array}{r} 1.008 \\ .927 \\ 1.097 \\ \hline \end{array}$ | NA | . 014 | A | . 018 |
|  |  | NA | .888- | NA | $\begin{aligned} & .961-060-1 \end{aligned}$ |  |
| NA Not applicable. * See table 2 for ratio descriptions. $\dagger$ See Text for design descriptions. Design I has 40 selections from 49 cells; Design II has 40 selections from 86 cells; Design III has 56 selections from 69 cells. <br> $\ddagger v_{y_{c s-m a n}^{2}}^{2}$ denotes relative variances for manually construc${ }^{2}$ cs-man ted controlled selection. |  |  |  |  |  |  |
|  |  |  |  |  |  |  |  |  |  |  |
| $\left[\begin{array}{lll} 5 & v^{2} & \text { denotes } \\ & y_{\text {Cs }} \\ & \text { cor } \\ \hline \end{array}\right.$ |  |  | relative variances for computer oenerater controlled selections. |  |  |  |

With relative variances from stratified random sampling as a base, in Tables 5 and 6 ordered systematic sampling and controlled selection are compared to the randomized design and then with each other. The construction of Tables 5 and 6 parallels that of Tables 3 and 4. In the North Central Region, only Design II of the three computer controlled selection designs is included. Table 5 contains the comparisons for aggregates. The results for ratios are given in Table 6. The variances are between psu relative variances, which are population values. No sampling occurs. The summary measures below the tables show that all column averages are below one, indicating variance reductions. Notice that the lowest figures are associated with controlled selection. The ranges related to controlled selection are wider but have lower bounds than those for ordered systematic selection. Outcomes might be different if research conditions were changed. But for the present investigation, controlled selection shows an advantage.

However, a few words of warning are in order. The observed relative variances are between psu variances. What the effect on total variance might be we are unable to say at this time. The within-psu component can vary widely for different

| $\begin{aligned} & \text { Vari- } \\ & \text { able } \\ & \text { Code } \\ & \text { Nos. * } \end{aligned}$ | North Central |  |  |  | South |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | $\begin{aligned} & v^{2}+ \\ & y_{\Delta r} \\ & (20 \text { strata, } \\ & 40 \text { sels. }) \end{aligned}$ | $\frac{v_{y_{s y s}^{2}}^{\ddagger}}{\frac{v^{2}}{y_{s i}^{+}}}$ | $\begin{aligned} & v_{y_{c o-c o m p}^{2}} \\ & \frac{v_{c}^{2}+}{y_{s i}^{+}} \end{aligned}$ | $\left\{\begin{array}{l} v^{2} \\ y_{c s}-\text { comp } \\ y^{2} \ddagger \\ y_{s y s} \end{array}\right.$ | $\left\{\begin{array}{l} v^{2}+ \\ y_{\Delta n} \\ (28 \text { strata, } \\ 56 \text { sels }) \end{array}\right.$ | $\begin{aligned} & v^{2} \ddagger \\ & y_{s y b} \\ & \frac{v_{s}^{2}}{y_{s i}+} \\ & \hline 7 \end{aligned}$ | $\left\lvert\, \begin{aligned} & v_{y_{c s}^{2}-c o m p}^{s} \\ & v_{y_{s x}^{2}}^{2} \end{aligned}\right.$ | $\left\{\begin{array}{l} v^{2} y_{c s-c o m p}^{s} \\ v_{c}^{2} \ddagger \\ y_{s u s} \end{array}\right.$ |
|  |  |  |  |  |  |  |  |  |
|  |  |  |  |  |  |  |  |  |
| 1 | 2 | 3 |  | 5 | 6 |  | - | 9 |
| 1 | . 000371 | . 972 | . 935 | . 962 | . 000558 | . 905 | . 949 | 1.049 |
| 2 | . 000520 | 968 | . 953 | . 984 | . 000759 | . 901 | . 980 | 1.087 |
| 3 | . 000766 | 1.027 | . 957 | . 932 | . 000940 | . 860 | . 905 | 1.052 |
| 4 | . 000776 | 1.004 | . 937 | . 933 | . 001228 | . 885 | . 896 | 1.013 |
| 5 | . 000928 | 972 | 849 | . 873 | . 003462 | . 824 | . 828 | 1.005 |
| 6 | . 002333 | . 912 | . 852 | . 935 | . 001707 | . 916 | . 825 | . 901 |
| 7 | . 042257 | 973 | 929 | . 955 | . 028347 | . 933 | . 997 | 1.068 |
| 8 | . 005042 | 937 | 840 | . 897 | . 011726 | . 878 | . 933 | 1.062 |
| 9 | . 014787 | . 930 | . 875 | . 941 | . 023536 | . 912 | . 960 | 1.053 |
| 10 | . 024203 | 1.001 | . 862 | . 861 | . 008975 | . 958 | . 697 | . 721 |
| 11 | . 042287 | . 926 | . 912 | . 984 | 143963 | . 818 | . 865 | 1.057 |
| 12 | . 000413 | . 911 | . 857 | . 941 | . 000898 | . 918 | . 948 | 1.033 |
| 13 | . 000318 | . 920 | 975 | 1.060 | . 000681 | . 912 | . 965 | 1.058 |
| 14 | . 003631 | . 973 | . 890 | . 915 | . 005724 | 876 | 889 | 1.014 |
| 15 | . 017755 | . 993 | . 868 | . 874 | . 005041 | . 897 | . 747 | . 833 |
| 16 | . 013538 | . 982 | . 704 | . 717 | . 003834 | . 944 | . 848 | . 899 |
| 17 | . 000449 | . 935 | . 971 | 1.039 | . 000774 | 901 | 947 | 1.052 |
| 18 | . 000298 | . 945 | . 971 | 1.028 | . 000702 | 898 | . 957 | 1.066 |
| 19 | . 000344 | . 955 | . 996 | 1.043 | . 000992 | . 906 | . 960 | 1.060 |
| 20 | 000866 | . 934 | . 904 | . 967 | . 001218 | . 918 | . 934 | 1.018 |
| 21 | . 000949 | 1.065 | . 906 | . 851 | . 001678 | .871 | . 896 | 1.029 |
| 22 | . 000685 | 947 | , 844 | . 890 | . 001470 | . 849 | . 882 | 1.039 |
| 23 | . 004228 | . 957 | . 945 | . 988 | . 009315 | . 889 | . 865 | . 973 |
| 24 | . 006346 | . 939 | . 796 | . 848 | . 014861 | . 917 | . 868 | . 947 |
| 25 | . 015520 | . 972 | . 679 | . 698 | . 048236 | . 961 | . 971 | 1.010 |
| 26 | . 034428 | . 988 | . 646 | . 655 | . 223552 | . 997 | . 928 | . 932 |
| 27 | . 003786 | . 899 | . 863 | . 959 | . 004657 | 1.017 | . 958 | . 942 |
| 28 | . 006711 | . 920 | . 962 | 1.045 | . 008695 | . 955 | . 873 | . 914 |
| 29 | . 199799 | 1.025 | . 746 | . 728 | . 077930 | . 965 | . 920 | . 953 |
| 30 | . 000962 | . 931 | . 773 | . 830 | . 000743 | . 935 | . 896 | . 958 |
| 31 | . 000749 | . 933 | . 920 | . 986 | . 000969 | . 905 | . 917 | 1.014 |
| 32 | . 000859 | . 914 | . 832 | . 910 | . 000866 | . 947 | . 876 | . 924 |
| Sumary Measures |  |  |  |  |  |  |  |  |
| Arithmetic means Ranges |  | . 958 | . 873 | 913 | NA | . 912 | . 902 | . 992 |
|  |  | $\begin{array}{r} .899 \\ 1.065 \\ \hline \end{array}$ | $\begin{array}{r} .646- \\ .996 \\ \hline \end{array}$ | $\begin{array}{r} .655 \\ 1.060 \\ \hline \end{array}$ | i $A^{\text {a }}$ | $\begin{array}{r} .824 \\ 1.017 \\ \hline \end{array}$ | $\begin{array}{r} .691- \\ .997 \\ \hline \end{array}$ | $\begin{array}{r} .721- \\ 1.087 \\ \hline \end{array}$ |
| NA Not applicable. * See Table 1 for variable descriptions. <br> $\pm V_{\#_{\Delta j}}^{2}$ denotes relative varfances for stratified random selections, two per stratum. <br> $\ddagger V^{2}$ denotes relative variances for ordered, systematic selections, two per stratum <br> ${ }^{4}$ bys $\quad(20$ strata in the North Central; 28 strata in the South $)$. <br> $5 \mathrm{~V}^{2}$ denotes relative variances for computer generated controlled selections ( 40 <br> ${ }^{4}$ cs-comb selections from 86 cells in the North Central; 56 selections from 69 <br> cells in the South). |  |  |  |  |  |  |  |  |
|  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |

populations and with different characteristics of the same population. Furthermore, when psu's selected with certainty are combined with noncertainty selections, the effect of the between psu variance component will be reduced.
COMPONENTS OF CONTROLLED SELECTION VARIANCES: The opportunity to study between and within pattern variance components is a dividend from controlled selection variance calculations as illustrated in Table 7 for selected ratios from manual and computer selections, Designs I and II in the North Central.

In Table 7, the pairs of variance components from the two designs illustrate that, for a given sample design and ratio, the within pattern variance components are constant no matter how many patterns are formed or by what means. Also notice that: (1) The within pattern component dominates the total variance; (2) A reduction in within pattern variances and an increase in the between components is the general result from the increased stratification in Design II. But the net result is a reduction in total variance.

Some optimal balance of within and between pattern variance components is needed to reduce both the within pattern components and the total variances. This is an activity for continuing research.
ESTIMATING SAMPLING VARTABILITY FROM A CONTROLLED SELECTION SAMPLE: There has been some reluctance to use controlled selection because it is not a measurable design. To obtain approximate measures of sampling variability from a single sample researchers may choose a technique designed for systematic sampling, a reasonable choice to make as the two sampling procedures are closely related.

| Variable <br> Code Nos. * | North Central |  |  |  | South |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | $\begin{aligned} & v_{y_{\Delta h}}^{+} \\ & (20 \text { strata, } \\ & 40 \text { sels. }) \end{aligned}$ | $\begin{gathered} \dot{v}_{y}^{2}{ }^{\dagger} \\ \frac{v_{\Delta y s}^{2}}{y_{\Delta r}^{\dagger}} \end{gathered}$ | $\begin{aligned} & v_{y_{c \Delta-c o m p}^{2}}^{s} \\ & \hline v_{y_{s n}^{2}}^{2} \\ & y_{s i} \end{aligned}$ |  | $\begin{array}{\|l} v^{2}+ \\ y_{\Delta n} \\ (28 \text { strata, } \\ 56 \text { sels. }) \end{array}$ | $\begin{aligned} & v_{y_{s y s}^{2}}^{\ddagger} \\ & \frac{v_{s}^{2} \dagger}{y_{s h}} \end{aligned}$ | $\frac{v^{2}{ }^{5}}{y_{c s-c o m p}}$ | $\begin{aligned} & v_{y_{c s}^{2}-c o m p}^{\varepsilon} \\ & \hline v_{y_{s y s}^{2}}^{\ddagger} \end{aligned}$ |
| 1 | 2 | 3 | 4 | 5 | 6 | 7 | 8 | 9 |
| 40 | . 000378 | 1.007 | . 733 | . 728 | . 000663 | . 851 | . 856 | 1.006 |
| 41 | . 024277 | . 996 | . 870 | . 873 | . 009510 | . 969 | . 736 | . 759 |
| 42 | . 001234 | 1.024 | . 824 | . 804 | . 002965 | . 819 | . 803 | . 980 |
| 43 | . 041759 | . 921 | . 909 | . 988 | . 143591 | . 820 | . 866 | 1.056 |
| 44 | . 039430 | . 964 | . 932 | . 967 | . 026825 | . 946 | 987 | 1.043 |
| 45 | . 000098 | . 966 | . 924 | . 956 | . 000120 | . 858 | . 789 | . 920 |
| 46 | . 000099 | . 910 | . 904 | . 993 | . 000079 | . 959 | 937 | . 978 |
| 47 | . 002969 | . 983 | . 389 | . 905 | . 003735 | . 874 | 898 | 1.028 |
| 48 | . 013450 | . 981 | . 729 | . 743 | . 004690 | . 967 | . 893 | . 924 |
| 49 | . 007492 | . 940 | 856 | . 910 | . 008289 | 945 | . 900 | . 953 |
| 50 | . 000098 | . 922 | . 949 | 1.029 | . 000023 | . 965 | . 915 | . 948 |
| 51 | . 000458 | . 955 | . 803 | . 841 | . 000616 | . 856 | . 860 | 1.006 |
| 52 | . 000398 | 1.106 | . 909 | . 822 | . 000393 | . 854 | . 877 | 1.027 |
| 53 | . 000647 | . 951 | . 752 | . 790 | . 001108 | . 879 | . 877 | . 998 |
| 54 | . 000042 | 1.017 | 849 | . 835 | . 000112 | . 897 | 818 | . 912 |
| 55 | . 001040 | . 920 | . 571 | . 621 | . 003584 | . 935 | . 974 | 1.042 |
| 56 | . 007472 | 1.042 | . 678 | . 651 | . 088799 | 1.006 | . 903 | . 898 |
| 57 | . 004415 | . 955 | . 952 | . 998 | . 006846 | . 993 | . 876 | . 882 |
| 58 | . 000843 | . 950 | . 780 | . 882 | . 000587 | . 883 | . 874 | . 990 |
| 59 | . 000122 | 1.015 | . 726 | . 715 | . 000123 | . 920 | . 901 | . 979 |
|  |  |  | Summe | ry | Measures |  |  |  |
| Arithmetic means |  | .976 | . 827 | . 850 | NA | .910 | . 877 | . 966 |
| Ranges |  | .910- | . $571-$ | .621- | NA | . 819 | .736- | . 759 - |
|  |  | 1.106 | . 952 | 1.029 |  | 1.006 | 987 | 1.056 |

[^0]Some investigations have been undertaken on this topic, but accomplishments are not yet sufficient for reporting.
ESTIMATES FOR DOMAINS: Another area for investigation is the precision of estimates of domains that are planned and included in controlled selection designs. Some researchers have suggested that controlled selection might be especially well suited for domain estimation. This suggestion is yet to be tested.

| Items | Design I |  | Design II |  |
| :---: | :---: | :---: | :---: | :---: |
|  | 41 manually constructed patterns | 29 computer constructed patterns | 47 manually constructed patterns | 34 computer constructed patterns |
| 1 | 2 | 3 | 4 | 5 |
| 42. Proportion of persons 65 years of age and over |  |  |  |  |
| Total |  |  |  |  |
| Var. (Y)* | 10,100 | 10,048 | 9,193 | 9,674 |
| Var. (X)* | 346,957 | 347,977 | 355,828 | 343,793 |
| Covar. ( $\mathrm{Y}, \mathrm{X}$ )* | 3,925 | 4,861 | 6,648 | 6,477 |
| Between patterns Var. (Y)* | 294 | 242 | 550 | 1,031 |
| Var. (X)* | 8,453 | 9,473 | 42,819 | 30,783 |
| Covar. ( $Y, X$ )* | -1,023 | -87 | -1,528 | -1,699 |
| Within patterns $\text { var. }(Y) *$ | 9,806 | 9,806 | 8,643 | 8,643 |
| Var. ( X )* | 338,504 | 338,504 | 313,010 | 313,010 |
| Covar. ( $\mathrm{Y}, \mathrm{X}$ ) * | 4,948 | 4,948 | 8,176 | 8,176 |
| 52. Proportion occupancy by movers into housing units during 1965-70 |  |  |  |  |
| Total |  |  |  |  |
| Var. (Y)* | 18,372 | 18,625 | 18,031 | 17.130 |
| Var. ( $X$ )* | 27,254 | 27,182 | 28,575 | 27,420 |
| Covar. ( $Y, X$ )* | 17,868 | 17,889 | 18,248 | 17,099 |
| Between patterns Var. ( $Y$ )* | 834 | 1,087 | 3,519 | 2,617 |
| Var. ( X ) \% | 671 | 599 | 3,504 | 2,349 |
| Covar. ( $\mathrm{Y}, \mathrm{X}$ ) * | 638 | 659 | 3,207 | 2,058 |
| Wethin patterns |  |  |  |  |
| Var. (Y)* | 17,539 | 17,539 | 14,513 | 14,513 |
| Var. ( X ) * | 26,583 | 26,583 | 25,071 | 25,071 |
| Covar. ( $\mathrm{Y}, \mathrm{X}$ ) * | 17,229 | 17,229 | 15,041 | 15,041 |
| 58. Average sales per retail trade establishment ( $\$ 1,000$ ) |  |  |  |  |
| Total |  |  |  |  |
| Var. (Y)* | 1,695,422 | 1,792,213 | 1,559,891 | 1,645,162 |
| Var. ( $X$ )* |  | 79 |  |  |
| Covar. ( $Y, X$ )* | 5,737 | 6,113 | 5,551 | 5,943 |
| Between patterns $\text { Var. }(Y) *$ | 70,974 | 167,765 | 153,427 | 238,698 |
| Var. ( X )* | 1 |  | 7 |  |
| Covar. ( $\mathrm{Y}, \mathrm{X}$ )* | 100 | 476 | 32 | 424 |
| within patterns Var. (Y)* | 1,624,448 | 1,624,448 | 1,406,464 | 1,406,464 |
| Var. (X)* | 75 | . 75 |  | 68 |
| Covar. ( $\mathrm{Y}, \mathrm{X}$ )* | 5,637 | 5,637 | 5,519 | 5,519 |

* In millions.

Appendix A
CONTROLLED SELECTION ILLUSTRATED: Simple examples of stratification for controlled selection are given in Illustration 1. In each example a population of 18 sampling units is assumed to have been distributed to 12 cells resulting from cross-tabulations by two variables, one with three and the other with four categories. While for simplification the numbers of variables and categories are kept small, each can be increased in practice. Also it is assumed that about 6 sample selections are to be made. The three examples illustrate design variations that achieve the desired sample size while distributing the sample across cells and marginals in proportion to their respective expectations.

In Example $X$, controlled selection can achieve a sample size of 5 with probability.3, or a sample of 6 with probability.7. Similar statements could be made for each cell.

In Example $y$, selection probabilities for the 18 sample units have been calculated to sum to exactly 6. Therefore, controlled selection would always yield a sample of 6 , with each of the marginals and cells tending to contribute pro-
portionately to the total sample according to the indicated probabilities. Row 1 would have 2 selections with probability .6 or 3 selections with probability .4 , and so on.

In Example $Z$, selection probabilities for individual units have been adjusted so that each row adds to exactiv 2. Adjustments could have been made to other integers, and in practice they are. Controlled selection would now achieve a sample of 6 with 2 selections from each row. Cells and columns would share proportionately in the sampie according to their respective expectations.

Illustration 2 displays a set of patterns or samples resulting from the application of controlled selection to the population described by Example Y. Pattern weights or values sum to 1.0 . Across all patterns the selection probabilities or expectations are satisfied exactly for every cell and marginal. Every pattern has 6 selections. Random choice of a pattern provides a sample that satisfies the specifications in Example $Y$.

Notice that controlled selection does not designate a particular set of sampling units. Controlled selection specifies the number of sampling units to be selected from designated cells. Illustrations 1 and 2 show that multiple selections from cells can occur and are permissible. The variance calculations assume that sampling units are chosen within cells in proportion to assigned probabilities and with replacement when multiple selections are to be made from a cell.


Illustration 2. Controlled Selection Patterns for Example Y Displayed in


Appendix B
ILLUSTRATION OF STRATIFIED DESIGNS WITH TWO SELECTIONS PER STRATUM: Two stratified designs are used in the research: stratified random, and ordered systematic. Simple examples of the selection designs are given in Illustrations 4 and 5, each derived from the strata in Illustration 3, which is an adaptation of Example $Y$ in Illustration 1, Appendix A.

The 18 sampling units are listed separately in a prescribed order that retains stratification by groups 1,2 and 3 , while reversing the order of classes $A$ through $D$ as the listing continues from one group to the next. The last class in group 1 is followed by the same class in group 2, and so on. Some of the unit probabilities were adjusted to force priba bilities to total 2 within each stratum. (While adjustments may appear to be major among the 18 primary units, in practice adjustments usually are within rounding error and affect only a few primary units.) With probabilities calculated tc one decimal, exactly ten pairs of primary units can be formed within a stratum, each pair receiving a weight of 0.1. (If psu probabilities were calculated to two or three decimals, 100 or 1,000 pairs would be formed in each stratum.) The method used to form pairs of units is the feature that distinguishes stratified random selection from the ordered systematic selection.

The stratified random selections displayed inIllustration 4 were made independently within each stratum by drawing 20 numbers at random to $f 111$ the ten cells. Numbers 01 and 11 were assigned to cell 1,02 and 12 to cell 2 , and so on. Notice that pSu Dl in stratum I had an adjusted probability of .4 in column 8, Illustration 3. Therefore, psu Dl appears four times (in positions 01, 04, 13 and 19) in stratum $I$, Illustration 4. Also notice that psuB2 with adjusted proba bility of .4 was assigned to positions $05,06,12$ and 15 . That is, in one cell B2 was paired with itself. Other selfpairings occurred in strata II and III. Utilizing a psu (in the selection process) according to its adjusted probability is inconsistent with random sampling with replacement, which could result in selecting a primary unit more or less frequently than its adjusted probability would indicate. However, the procedure that was used was preferred for the purpose of calculating population variances at a later stage in the research.

Ordered systematic selection restricts the combinations of sampling units that enter a particular sample. Judicious ordering of strata and sampling units are important steps in ordered systematic selection. The arrangement of strata and units in Illustration 3 was designed to satisfy the systeIllustration 3. Equal Sized Strata Constructed from Data in Example $Y$

| Strata | Groups | Classes | Sums of <br> Proba- <br> bili- <br> ties <br> for <br> Cells | $\begin{aligned} & \text { PSU } \\ & \text { codes } \end{aligned}$ | PSU <br> Proba- <br> bili- <br> ties | Cumula- <br> tive <br> Sums of Proba- <br> bili- <br> ties | Adjusted Probabilities* |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  |  |  |  |  | $\begin{aligned} & \text { For } \\ & \text { PSU's } \end{aligned}$ | Cumula <br> tive <br> Sums |
| T | 2 | 3 | 4 | 5 | 6 | 7 | 8 | 9 |
| I | 1 | D | 1.5 | 1 | . 5 | . 5 | . 4 | . 4 |
|  |  |  |  | 2 | . 4 | . 9 | . 4 | . 8 |
|  |  |  |  | 3 | . 3 | 1.2 | . 3 | 1.1 |
|  |  |  |  | 4 | . 3 | 1.5 | . 2 | 1.3 |
|  |  | B | . 7 | 1 | . 3 | 1.8 | . 3 | 1.6 |
|  |  |  |  | 2 | . 4 | 2.2 | . 4 | 2.0 |
| II | 1 | A | . 2 | 1 | . 2 | 2.4 | . 2 | 2.2 |
|  |  | A | . 5 | 2 | . 5 | 2.9 | . 6 | 2.8 |
|  |  | B | . 3 | 1 | . 3 | 3.2 | . 4 | 3.2 |
|  |  | C | . 4 | 1 | . 4 | 3.6 | . 5 | 3.7 |
|  |  | D | . 2 | 1 | . 2 | 3.8 | . 3 | 4.0 |
| III | 3 | c | . 6 | 1 | . 2 | 4.0 | . 2 | 4.2 |
|  |  |  |  | 2 | . 4 | 4.4 | . 4 | 4.6 |
|  |  | B | . 2 | 1 | . 2 | 4.6 | . 2 | 4.8 |
|  |  | A | 1.4 | 1 | . 5 | 5.1 | . 4 | 5.2 |
|  |  |  |  | 2 | . 4 | 5.5 | . 4 | 5.6 |
|  |  |  |  | 3 | . 3 | 5.8 | . 2 | 5.8 |
|  |  |  |  | 4 | . 2 | 6.0 | . 2 | 6.0 |

[^1]matic design, and that ordering was maintained when forming the pairs of psu's shown in Illustration 5 .

Notice the difference between the assignment of primary untis to cells in Illustrations 4 and 5. In stratum $I$ of Illustration 5, psu D1 is assigned to four positions in sequence, 01, 02,03, and 04. Then psu D2 is assigned to the next four positions. Next psu's D3, D4, B1 and B2 follow in sequence and in accordance with their adjusted probabilities. There are no cases of self-pairing of psu's in illustration 5, nor would there be unless unusual conditions prevailed. (One psu might have probability greater than .5.) The ordered assignment of psu's to cell positions continues from stratum I into II and throughout stratum III.

Illustration 4. A Stratif ted Random Design with Two Selections per Strazum, Chosen in Proportion to Asbigned Probablll:tes and with Replacement*

| Strata | Sample Numbers |  |  |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | 1 | 2 | 3 | 4 | 5 | 6 | 7 | 8 | 9 | 10 |
|  | Heights |  |  |  |  |  |  |  |  |  |
|  | .1 | .1 | . 1 | . 1 | . 1 | . 1 | . 1 | . 1 | . 1 | . 1 |
| I | DI | D2 | D2 | D1 | 32 | 82 | D3 | D2 | $B 1$ | 04 |
|  | 04 | B2 | 01 | D3 | 82 | 81 | D2 | 81 | 01 | D3 |
| II | B1 | A2 | A2 | C) | $B 1$ | Cl | 81 | ${ }^{4} 2$ | 01 | A2 |
|  | A1 | A2 | A1 | C) | Cl | 01 | 01 | D 1 | A2 | B1 |
| 111 | C2 | A2 | A2 | A4 | A 4 | ${ }_{6} 1$ | ${ }^{1}$ | C2 | 81 | ${ }^{\text {B1 }}$ |
|  | 42 | A 3 | A 1 | Al | Al | C2 | A2 | A3 | A2 | Al |

* Notice that the number of appearances of each psu agrees exactly with its adjusted unrestricted random selections, the lllustrated formation, a convenient device for variance calculations, was used in calculating the variances shown in Tables 5 and 6

Illustration 5. An Ordered Systematis Design with Two Selections per Stratum*

| Strata | Sample Numbers |  |  |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | 1 | 2 | 3 | 4 | 5 | 6 | 7 | 8 | 9 | 10 |
|  | Weights |  |  |  |  |  |  |  |  |  |
|  | . 1 | 1 | .1 | 1 | 1 |  | 1 | 1 | 1 | 1 |
| I | D1 | D1 | 01 | 01 | 02 | 02 | D2 | D2 | D3 | D3 |
|  |  | 04 | 04 | 81 | B1 | 81 | B2 | 82 | 82 | B2 |
| II | A) | Al | A2 | A2 | A2 | A2 | A2 | A2 | 81 | B1 |
|  |  | B1 | $\mathrm{Cl}^{1}$ | 4 | Cl | 01 | CI | D1 | 01 | D1 |
| III | C1 | C1 | C2 | c2 | C2 | C2 | 81 | 日 1 | A1 | A1 |
|  | A1 | Al | A2 | A2 | A2 | A2 | A3 | A3 | A4 | A 4 |

of lllustration 3 .
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Michigan 48106 .


[^0]:    $+v_{y_{3 \mu}}^{2}$ denotes relative variances for stratified random selections, two per stratum. $\mp v_{y_{s / \pi}}$ $y_{\text {sus }}^{2}$ denotes relative variances for ordered, systematic selections, two per stratum ${ }^{5} \mathrm{~V}^{2}$ (20 strata in the North Central; 28 strata in the South).
    $y_{c o-c}$ denotes relative variances for computer generated controlled selections (40 $y_{\text {co-comp }} \begin{aligned} & \text { selections from } 86 \text { cells in the North Central; } 56 \text { selections from } 69 \\ & \text { cells in the South). }\end{aligned}$

[^1]:    * In practice, adjusted probabilities ususally will be within rounding error of initially calculated probabilities.

